

# Two-Stage Designs for proteomic and gene expression studies applying methods differing in costs

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# Motivation

In genomic or proteomic studies we generally search for genes or proteins which explain the occurrence of a disease, the cause of a disease, ...

## Limitations

- very large number of candidate genes (proteins)
- only a very small fraction of genes (proteins) show noticeable effects

## Constraints

- Limit of resources: total costs  $C$  are fixed

# Test Problem:

## Consider

$m_1$  hypotheses for the mean of independent normally distributed observations  $\mu_i$  with known variance  $\sigma^2$

$$H_{0i} : \mu_i = 0 \quad \text{versus} \quad H_{1i} : \mu_i > 0 \quad i = 1, \dots, m_1$$

assuming independence of observations across hypotheses

# Motivation: Example

## Consider an experiment with

- $m_1 = 1000$  ... number of hypotheses tests
- $C = 20000$  ... fixed total costs

## Given

- $\pi_0 = 0.99$  ... proportion of true null hypotheses among all  $m_1$  hypotheses
- $\Delta = \mu/\sigma = 0.75$  ... effect size

## Conventional Single-Stage Design

- The costs per observation are set to  $c_1 = 1$ .
- Distribute total costs  $C = 20000$  equally among the hypotheses  
 $\Rightarrow 20000/1000 = 20$  observations per hypothesis test

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# Error Rates

## Notation

$R$  ... number of rejected hypotheses

$FP$  ... number of false positive decisions

## Family Wise Type I Error Rate (FWE)

is defined as the probability of at least one Type I error:

$$FWER = P(FP \geq 1)$$

## False Discovery Rate (FDR)

is the expected proportion of Type I errors among the rejected hypotheses (Benjamini and Hochberg, 1995):

$$FDR = E \left( \frac{FP}{\max\{R, 1\}} \right)$$

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# Controlling the Family Wise Type I Error Rate

## Bonferroni adjusted level

- Reject all hypotheses with a p-value  $p_i \leq \gamma$
- $\gamma = \alpha/m_1$

## Example: Single-Stage Design with 20 observations

$$\alpha = 0.05: \quad \gamma = 0.05/1000 = 0.00005$$

Assuming an effect size of  $\Delta = 0.75$  the Power (expected fraction of correctly identified effective genes/proteins) is **0.296**

Because we have to adjust for multiplicity, the power of the single-stage design is very low as compared to a single hypothesis test with the same sample size.

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# Controlling the False Discovery Rate (Storey 2002)

## Storey's procedure to control the FDR at level $\alpha$

- $m_1$  hypothesis tests
- Reject all hypotheses with a p-value  $p_i \leq \gamma$
- To control the FDR at level  $\alpha$  choose the maximal  $\gamma$  so that

$$\widehat{FDR}_\gamma(p_1, \dots, p_{m_1}) = \frac{\gamma \hat{\pi}_0 m_1}{\max(\#\{p_i \leq \gamma\}, 1)} \leq \alpha$$

where  $\hat{\pi}_0 = \#\{p_i > \lambda\} / \{(1 - \lambda)m_1\}$  and  $0 < \lambda < 1$

## Example: Single-Stage Design with 20 observations

For  $\Delta = 0.75$ :    *Power* = 0.443    (FWE: *Power* = 0.296)

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## Single-stage designs

Single-stage designs have low power to detect existing effects.

## Two-stage designs

Screen for promising hypotheses in the first stage which are further investigated in the second stage:

- Limit of resources: total costs  $C$  are fixed
- A fraction  $r$  of the resources  $C$  are used in the first stage for screening
- The remaining resources  $(1 - r)C$  are used for second stage
- Costs per observation are set to 1

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# The Two-Stage Design

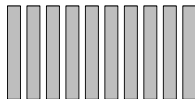
$$n_1 = \frac{rC}{m_1}$$

$m_1$  hypotheses



Select  $H_i: p_{i1} < \gamma_1$

$m_2$  (random)



$$n_2 = \frac{(1-r)C}{m_2}$$

Test decision: Reject  $H_i: p_{i2} < \gamma_2$

The computation of the p-value after the second stage depends on the ...

...design of the two-stage experiment

- Pilot Design
- Integrated Design

# The Pilot Design

$m_1$  hypotheses

$n_1$

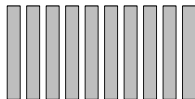


}  $p_{i1}$

Select  $H_j: p_{i1} < \gamma_1$

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$n_2$



}  $p_{i2}$

Test decision: Reject  $H_j: p_{i2} < \gamma_2$

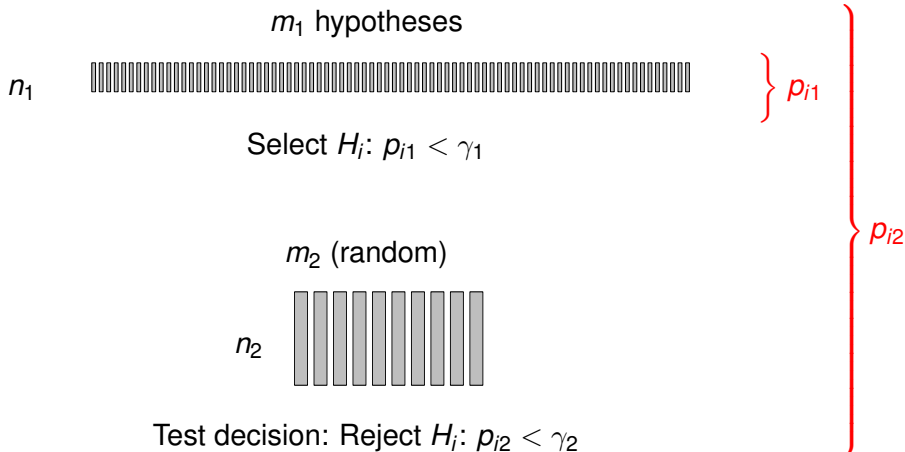
# The Pilot Design

The final test decision after the second stage is only based on the second stage data:

→ Sample size for the  $m_2$  final hypothesis tests is  $n_2$ .

The correction for multiplicity is performed for the  $m_2$  selected hypotheses using Bonferroni's correction (FWE) or Storey's procedure (FDR).

# The Integrated Design



# The Integrated Design

The final test decision is based on the pooled observations over both stages:

- Sample size for  $m_2$  final hypothesis tests is  $n_1 + n_2$ .
- Price to be paid: the correction has to be performed for all initial  $m_1$  hypotheses.

## Error Rate Control

- early acceptance decision in a sequential test procedure
- define sequential p-values for all  $m_1$  hypotheses on the Tsiatis-Rosner-Metha (1984) ordering in the sample space:  
For hypotheses early accepted we use the first stage p-value, for the final analysis we integrate over all sample path more extreme than the observed one.
- Despite that  $m_2$  and  $n_2$  are random the sequential p-values are independent and uniformly distributed under the null hypothesis.  
(see Zehetmayer et al, 2005)

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# Asymptotic optimal designs

## Two-stage designs

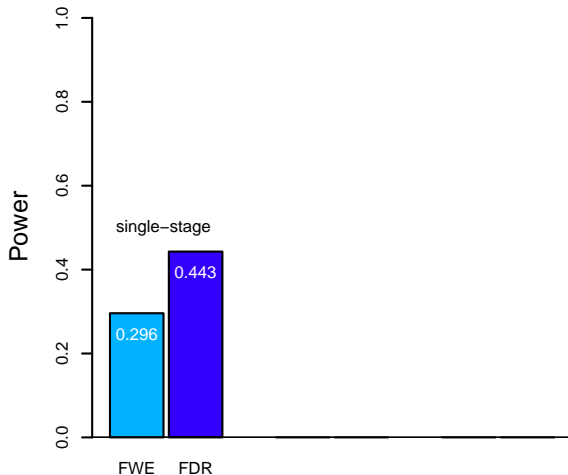
For given overall costs  $C$ , proportion of true null hypotheses  $\pi_0$  and effect size  $\Delta = \mu/\sigma$  the power of a two-stage design can be optimized with respect to:

- $r$  ..fraction of total costs used in the first stage
- $\gamma_1$  ..selection boundary after first stage

Asymptotic: for  $m_1 \rightarrow \infty$  and  $C = d * m_1$  for  $d > 0$

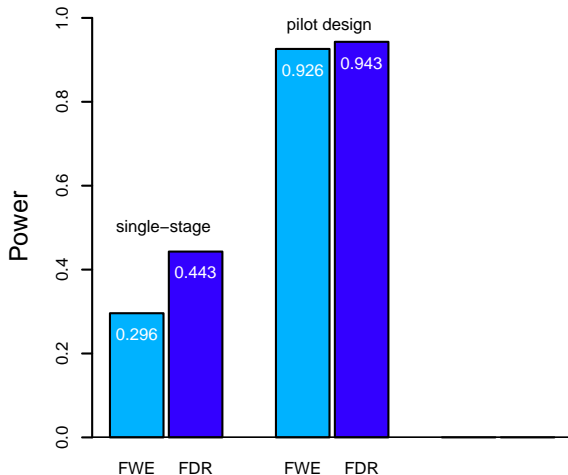
# Asymptotic optimal designs

$C = 20000$ ,  $m_1 = 1000$ ,  $\Delta = 0.75$ ,  $\alpha = 0.05$ ,  $\pi_0 = 0.99$



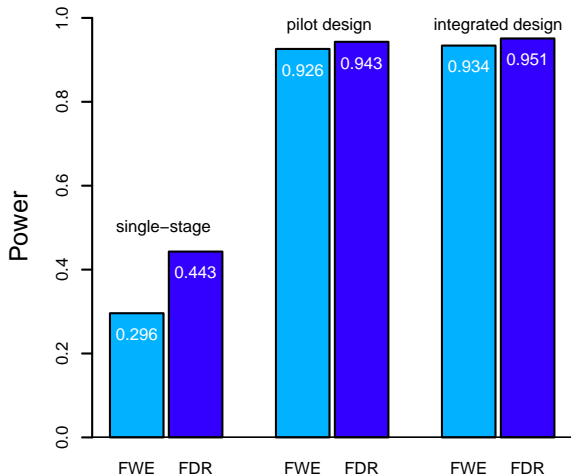
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# Example:

## Pilot Design controlling the FWE

- 71.8 % of  $C$  are used in the first stage
- $n_1 = 20000 * 0.718/1000 \approx 14$
- $\gamma_1 = 0.092$
- $m_2 \approx 100$  hypotheses are selected for the second stage
- $\gamma_2 \approx 0.05/100 = 0.0005$
- $n_2 = (1 - 0.718) * 20000/100 \approx 56$

# In Genomic or Proteomic studies ...

## Scenario 1: Different costs

Different costs per observation may arise at both stages:

- costs per observation in the first stage set to  $c_1 = 1$
- cost ratio between stages  $c_2 > 1$

## Scenario 2: Different costs and effect sizes

There is an increasing focus on using a less accurate assay in early stages and a more accurate one in later stages:

- cost ratio  $c_2 > 1$
- effect size ratio between stages  $k > 1$

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# The Two-Stage Design: Scenario 1

$$n_1 = \frac{rC}{m_1}$$

$m_1$  hypotheses

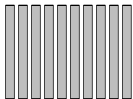


effect size



Select  $H_j: p_{j1} < \gamma_1$

$m_2$  (random)



effect size



$$n_2 = \frac{(1-r)C}{c_2 m_2}$$

Test decision: Reject  $H_j: p_{j2} < \gamma_2$

# The Two-Stage Design: Scenario 2

$$n_1 = \frac{rC}{m_1}$$

$m_1$  hypotheses

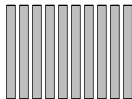


effect size



Select  $H_j: p_{j1} < \gamma_1$

$m_2$  (random)



effect size



$$n_2 = \frac{(1-r)C}{c_2 m_2}$$

Test decision: Reject  $H_j: p_{j2} < \gamma_2$

# Scenario 1: Example

Consider again the experiment with

- $m_1 = 1000$  ... number of hypotheses tests
- $C = 20000$  ... fixed total costs (limit of resources)

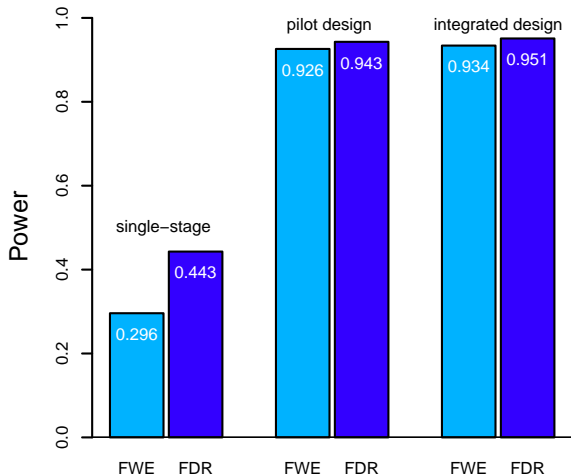
Given

- $\pi_0 = 0.99$  ... proportion of true null hypotheses among all  $m_1$
- $\Delta = 0.75$  ... effect size

We now assume that there a cost ratio between stages  $c_2 > 1$ .

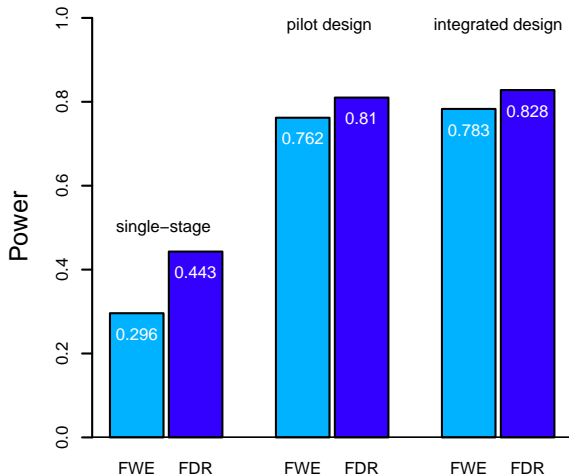
# Asymptotic optimal designs

$C = 20000$ ,  $m_1 = 1000$ ,  $\Delta = 0.75$ ,  $\alpha = 0.05$ ,  $\pi_0 = 0.99$ ,  $c_2 = 1$



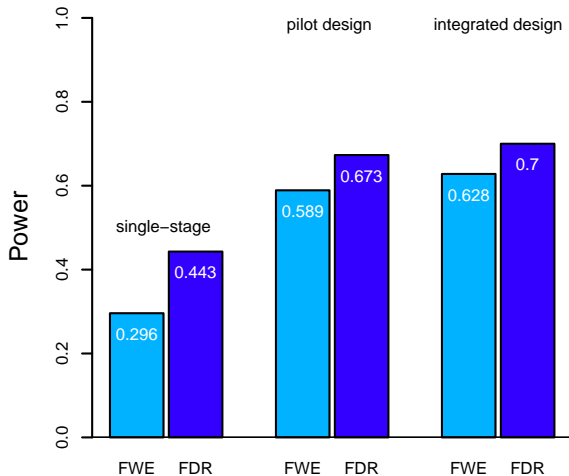
# Asymptotic optimal designs

$C = 20000$ ,  $m_1 = 1000$ ,  $\Delta = 0.75$ ,  $\alpha = 0.05$ ,  $\pi_0 = 0.99$ ,  $c_2 = 5$



# Asymptotic optimal designs

$C = 20000$ ,  $m_1 = 1000$ ,  $\Delta = 0.75$ ,  $\alpha = 0.05$ ,  $\pi_0 = 0.99$ ,  $c_2 = 15$



# Break Even Point in Cost Ratio

If the cost ratio  $c_2$  increases, the power of the two-stage design decreases.

## Question

Is there a cost ratio  $c_2^*$ , where it does not make sense to apply a two-stage design as compared to the single-stage design?

## Integrated design

$c_2^*$  does not exist:

The power of the asymptotic optimal integrated design

- is always larger than of the corresponding single-stage design.
- converges to the power of the single-stage design ( $\lim_{c_2 \rightarrow \infty} r = 1$ )

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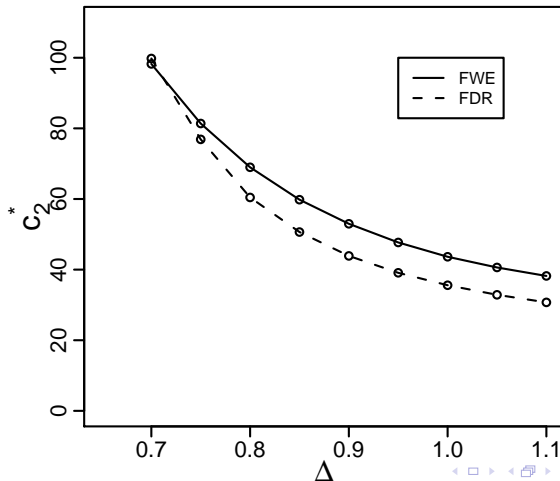
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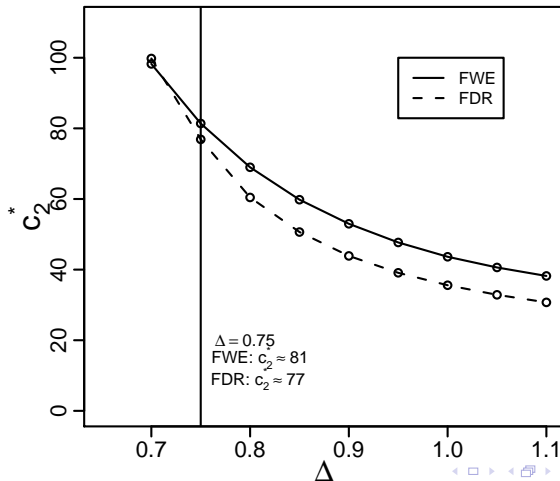
# Break Even Point: Pilot Design

$C = 20000$ ,  $m_1 = 1000$ ,  $\pi_0 = 0.99$ ,  $\alpha = 0.05$



# Break Even Point: Pilot Design

$$C = 20000, m_1 = 1000, \pi_0 = 0.99, \alpha = 0.05$$



# Break Even Point in Cost Ratio

- For the integrated design no such Break Even Point exists.
- For the pilot design such point exist but for small  $\Delta$  it is large.
- Two-stage designs are a good option to increase the power even if the cost ratio per observation between stages is fairly high.

# Impact of Design Misspecifications

Whereas costs are usually known a priori the optimal designs depend on the unknown parameters  $\pi_0$  and  $\Delta$ .

Is there an amount of misspecification where it would have been better to use a single-stage design?

## Example

- $C = 20000$
- $m_1 = 1000$
- $c_2 = 15$

$r$  and  $\gamma_1$  planned for the situation:

- $\pi_0 = 0.99$
- $\Delta = 0.75$

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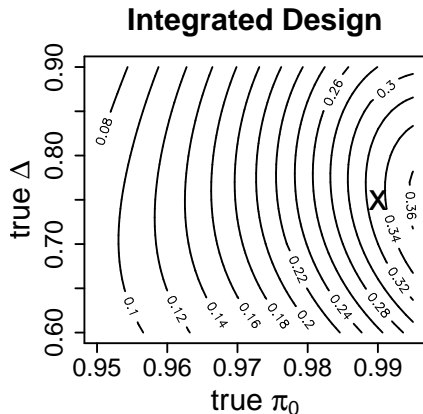
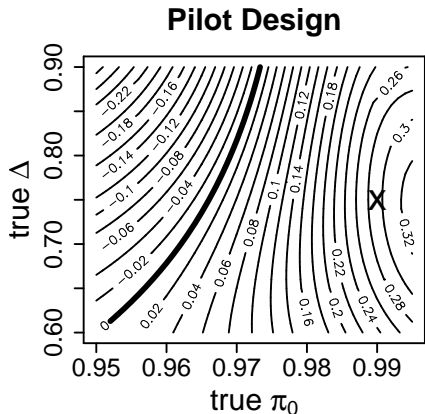
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- $\pi_0 = 0.99$
- $\Delta = 0.75$

Difference of power values between the two-stage designs and the corresponding single-stage designs (Control of FWE):



X ... point for which two-stage design parameters are optimized

# Misspecifications

- The integrated design is more robust against design misspecifications
- if the planned  $\pi_0$  is larger than the true one:  
loss of power as compared to the single-stage design
- if the planned  $\pi_0$  is smaller than the true one:  
increase of power as compared to the single-stage design

## Scenario 2

The experimenter has two different candidate methods for the measurements from the very beginning:

- low-cost standard method: effect size=  $\Delta$   
costs per observation = 1
- high-cost improved method: effect size=  $k\Delta$   
costs per observation  $> 1$
- cost ratio between methods  $c_2 > 1$
- effect size ratio between methods  $k > 1$

### Two-Stage Procedures

- first stage: low      second stage: low
- first stage: low      second stage: high
- first stage: high      second stage: high

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## Two-Stage Procedures

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- first stage: low      second stage: high
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# Pilot Design controlling the FWE

Consider an experiment with

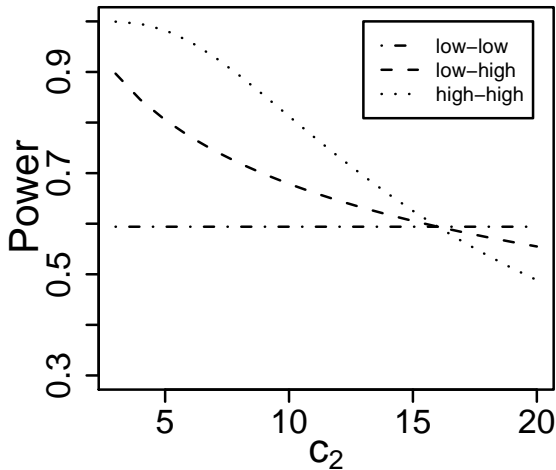
- $\alpha = 0.05$
- $C = 20000, m_1 = 1000$
- low-cost method: effect size  $\Delta = 0.5$
- high-cost method:  $k = 4$  effect size:  $0.5 * 4 = 2$

## Question

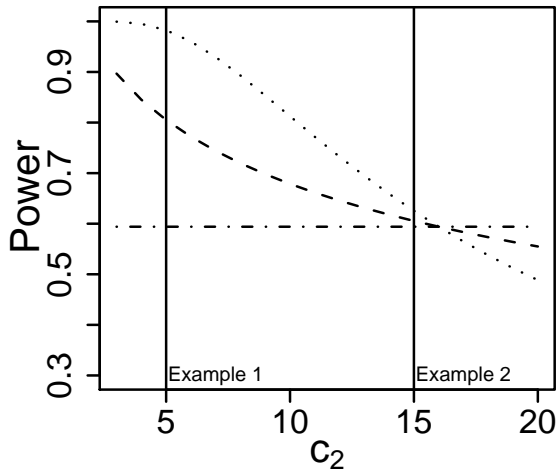
Looking at different values for  $c_2$ :

Which of the 3 procedures has the maximal Power?

# Common crossing point at $c_2 = k^2 = 16$



# Common crossing point at $c_2 = k^2 = 16$



# Examples:

Procedure		Example 1: $c_2 = 5$	Example 2: $c_2 = 15$
low-low	Power	0.594	0.594
	$n_1$	$\approx 13$	$\approx 13$
low-high	Power	0.805	0.605
	$n_1$	$\approx 10$	$\approx 13$
high-high	Power	0.983	0.625
	$n_1$	$\approx 3$	$< 1$

## Scenario 2

- The low-high procedure may only be preferable if the high-cost method is too expensive to achieve the necessary repetitions at the first stage.
- In our example the high-cost method is not too expensive but highly effective. If we have only **double** effects for the expensive method than the low-low procedure would be already preferable if the costs are only **4** times larger.
- If the improved method is much more expensive it has to be much more effective to apply a high-high or a low-high procedure.

# Conclusions

## Different costs in both stages

Two-stage designs are a good option to improve the power even if the cost ratio between stages  $c_2$  is fairly high.








## Misspecification

- The integrated design is more robust against misspecification than the pilot design.
- Optimism in the planning phase with regard to the number of true alternatives may help to avoid loss of power.

## If two different methods are available

- Depending on  $c_2$  and  $k$  it is preferable to run two-stage designs which apply either the low-cost or the high-cost method at both stages.
- Switching from the low-cost to the high-cost method may only be advisable if there is lack of finance so that  $n_1$  for the high-cost method is too small.

# Selected References

-  Benjamini,Y, Hochberg,Y (1995) Controlling the false discovery rate: a practical and powerful approach to multiple testing.  
*J. R. Statist. Soc. B*, 57, 289-300.
-  Goll,A, Bauer,P (2007) Two-stage designs applying methods differing in costs.  
*Bioinformatics*, 23, 1519-1526.
-  Satagopan,JM et al. (2004) Two-stage designs for gene-disease association studies. *Biometrics*, 58, 163-170.
-  Storey,JD (2002) A direct approach to false discovery rates.  
*J. R. Statist. Soc. B*, 64, 479-498.
-  Storey,JD, Taylor,JE, Siegmund,D (2004) Strong control, conservative point estimation and simultaneous conservative consistency of false discovery rates: a unified approach. *J. R. Statist. Soc. B*, 66, 187-205.
-  Wang,H et al. (2006) Optimal two-stage genotyping designs for genome-wide association scans, *Genetic Epidemiology*, 30, 356-368.
-  Zehetmayer,S, Bauer,P, Posch,M (2005) Two-stage designs for experiments with a large number of hypotheses.  
*Bioinformatics*, 21, 3771-3777.